

## Probability Theory Refresher (§2)

Will be slightly informal (but in a way that can be made completely rigorous)  
 Axiomatic approach → text book / after class. When in doubt: ASK!

**Probability distribution** on  $\mathcal{A}$  (finite set):  $P: \mathcal{A} \rightarrow \mathbb{R}_{\geq 0}, \sum_{a \in \mathcal{A}} P(a) = 1$

e.g. Bernoulli( $f$ ):  $\mathcal{A} = \{0, 1\}, P(1) = f, P(0) = 1-f$

Uniform( $\mathcal{A}$ ):  $P(a) = \frac{1}{|\mathcal{A}|} \quad \forall a \in \mathcal{A}$

**Random variable (RV)**  $X \stackrel{\Delta}{=} \text{prob. dist. } P_X \text{ on set } \mathcal{A}_X$

**NOTATION:**  $X \sim P$  for  $P_X = P$  leave out subscript & it's clear

UNLIKE THE  
BOOK, I ALWAYS  
DISTINGUISH  
 $X$  and  $x$

$$\Pr(X=x) = P_X(x) \stackrel{\text{def}}{=} P(x)$$

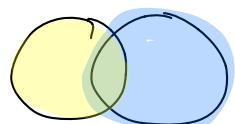
$$\Pr(X \in S) = \sum_{x \in S} P(x)$$

$$\Pr(\text{condition on } X) = \sum_{x \text{ condition holds}} P(x) = \Pr(X \in \{x \text{ s.t. condition holds}\})$$

e.g. if  $X$  random variable on  $\{1, \dots, 6\}$ :

$$\Pr(X \text{ even}, X \neq 2) = \Pr(X \in \{4, 6\}) = P(4) + P(6)$$

$$* \Pr(A \text{ or } B) = \Pr(A) + \Pr(B) - \Pr(A \text{ and } B) \quad \begin{array}{l} \text{if mutually exclusive} \\ \Pr(A) + \Pr(B) \\ \text{"union bound"} \end{array}$$



\*  $X$  RV,  $f$  function  $\Rightarrow Y = f(X)$  RV

$$\Pr(Y=y) = \sum_{x: f(x)=y} \Pr(X=x) \quad \text{or simply} \quad P(y) = \sum_{f(x)=y} P(x)$$

More than one random variable

How to describe "pair of RVs"  $(X, Y)$ ? "Joint" prob. dist.:

$$\Pr(X=x, Y=y) = P_{(X,Y)}(x,y) = P_{XY}(x,y) = P(x,y)$$

i.e.  $(X, Y)$  is RV on  $\mathcal{A}_{XY} = \mathcal{A}_X \times \mathcal{A}_Y$ . Similar for tuples.

\* Can visualize by "contingency table":

\* marginal distributions of  $X$  &  $Y$ :

$$P(X) = \sum_Y P(x,y) \quad \& \quad P(Y) = \sum_X P(x,y)$$

i.e.

$$\Pr(X=x) = \sum_Y \Pr(X=x, Y=y) \quad \text{etc.}$$

$Y \setminus X$	SUMMER	WINTER	
SUN	30%	10%	40%
RAIN	20%	40%	60%
	50%	50%	

NOT independent!  
 $P(\text{SUN, SUMMER}) \neq P(\text{SUN}) \cdot P(\text{SUMMER})$

\*  $X, Y$  are called independent if  $P(x,y) = P(x) \cdot P(y)$

How about:

Conditional prob. dist. of  $Y$  given  $X$ :

$$\Pr(Y=y | X=x) := \frac{\Pr(X=x, Y=y)}{\Pr(X=x)}$$

15%	60%
5%	20%

NOTATION:  $P_{Y|X=x}(y), P_{Y|X}(y|x), P(Y|x), \dots$

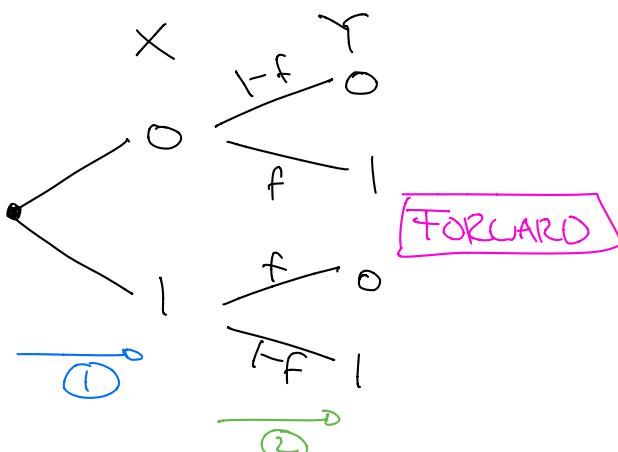
i.e.  $P(Y|x) = \frac{P(x,y)}{P(x)}$  and  $P(x|y) = \frac{P(x,y)}{P(y)}$

\*  $P(Y|x)$  is prob. dist in  $y$  for each fixed  $x$   $\in \mathcal{X}$

Two simple rewritings:

$$* P(x|y) = \underset{(1)}{P(x)} \underset{(2)}{P(y|x)} = P(Y) P(x|Y)$$

e.g.  $X$  channel input,  $P(Y|X)$  channel  
 $Y$  channel output



\* Bayes rule:

$$P(x|y) = \frac{P(x|y) P(y)}{P(x)} = \frac{P(x|y) P(y)}{\sum_{y'} P(x|y') P(y')}$$

INVERSE

e.g.  $P(\text{pos} | \text{sick}) = P(\text{neg} | \text{healthy}) = 90\%, P(\text{sick}) = 1\%$

$$\Rightarrow P(\text{sick} | \text{pos}) = \frac{0.9 \times 0.01}{0.9 \times 0.01 + 0.1 \times 0.99} = \frac{1}{12} < 10\% \quad ?$$

E.g. decoding the repetition code  $R_3$ : assume  $S \sim \text{Uniform}(\{0,1\})$  all indep.  
 $R_1 = S \oplus N_1, \dots, R_3 = S \oplus N_3$  were  $N_1, N_2, N_3 \sim \text{Bern}(f)$

Assume we received  $r = r_1 r_2 r_3$ . How should we estimate  $S$ ?

$$P(S|r) = \frac{P(r|S) P(S)}{P(r)} \underset{\text{fixed}}{=} \frac{P(r|S=0)}{P(r|S=1)} = \frac{P(r|T=000)}{P(r|T=111)} = \prod_{k=1}^3 \frac{P(r_k | T_k = 0)}{P(r_k | T_k = 1)}$$

$$\Rightarrow \frac{P(S=0|r)}{P(S=1|r)} = \frac{P(r|S=0)}{P(r|S=1)} = \begin{cases} > 1 & \text{if } \#0's > \#1's \\ < 1 & \text{if } \#1's > \#0's \end{cases} = \text{majority vote}$$

$$= \left(\frac{1-f}{f}\right)^{\#0's} \cdot \left(\frac{f}{1-f}\right)^{\#1's} \underset{> 1 \text{ since } f < 50\%}{=}$$

$$\frac{1-f}{f} \text{ if } r_k = 0, \text{ else } \frac{f}{1-f}$$

Combining independent RV's:

Ques: ① Let  $S, k \sim \text{Uniform}(\{0,1\})$ ,  $T = S \oplus k$ . Yes:  $P(S=s, T=t) = P(S=s, k=s \oplus t) = \frac{1}{4}$   
 Are  $S$  and  $T$  independent?

② How to label two dice w/ numbers from  $\{0, 1, \dots, 6\}$  s.t. their sum  $\sim \text{Uniform}(\{1, 2, \dots, 12\})$

A: 123456  
 B: 000666

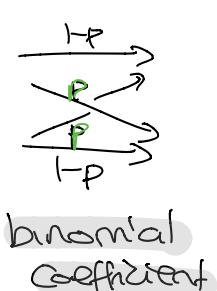
Binomial( $n, p$ ): Distribution of  $Y = X_1 + \dots + X_n$  were  $X_i \stackrel{iid}{\sim} \text{Bernoulli}(p)$

\* e.g. number of bit flips when we send  $n$  bits through

$$\Pr(Y=k) = \binom{n}{k} p^k (1-p)^{n-k}$$

$\underbrace{\phantom{\Pr(Y=k) = \binom{n}{k} p^k (1-p)^{n-k}}}_{\substack{\# \text{bitstrings} \\ \text{with } k \text{ ones} \\ n-k \text{ zeros}}} \quad \underbrace{\phantom{\Pr(Y=k) = \binom{n}{k} p^k (1-p)^{n-k}}}_{\substack{\text{probability} \\ \text{of any such} \\ \text{string}}}$

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$



"Numerical" random variables

If  $X \sim P$  is RV with values in  $c \subseteq \mathbb{R}$ :

Expectation value (mean):  $E[X] = \sum_x P(x) \cdot x$

\*  $E[f(X)] = \sum_x P(x) \cdot f(x)$  "law of the unconscious statistician"

\*  $E[cX] = c \cdot E[X]$  &  $E[X+Y] = E[X] + E[Y]$  (A)

\* If  $X, Y$  independent:  $E[XY] = E[X] \cdot E[Y] = \sum_{x,y} p(x)p(y)xy$

$\hookrightarrow X \sim \text{Uniform}(\{-1, 1\}), Y = -X$   $\stackrel{\text{NOT indep}}{\Rightarrow} E[XY] = -1, E[X] = E[Y] = 0$

Variance:  $\text{Var}(X) = E[(X - EX)^2]$

$$= \sum_x P(x)(x - EX)^2 = E[X^2] - E[X]^2$$

\*  $\text{Var}(cX) = c^2 \text{Var}(X)$

\* If  $X, Y$  independent:

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) \quad (\text{B})$$

$\vdash$  use that  $E[XY] = E[X] \cdot E[Y]$

### Examples

P	Bernoulli( $p$ )	Binomial( $n, p$ )
E	$p$	$n \cdot p$ (A)
Var	$p(1-p)$	$n \cdot p \cdot (1-p)$ (B)

$$\begin{aligned} & p \cdot (1-p)^2 + (1-p) \cdot (0-p)^2 \\ & = p(1-p) \end{aligned}$$

Interpretation?

Markov inequality: If  $X \geq 0$ :  $\Pr(X \geq t) \leq \frac{E[X]}{t} \quad (\forall t > 0)$

PF:  $\Pr(X \geq t) = \sum_{x \geq t} P(x) \leq \sum_{x \geq t} P(x) \cdot \frac{x}{t} \leq \frac{E[X]}{t} \quad \square$

Chebyshev inequality:  $\Pr(|X - EX| \geq \varepsilon) \leq \frac{\text{Var}(X)}{\varepsilon^2}$  WHP deviation from mean is of order  $\sqrt{\text{Var}(X)}$

PF: Apply Markov to  $Y = (X - EX)^2$ . □

Law of large numbers:  $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} P$  w/ mean  $\mu$ , variance  $\sigma^2$ ,

$$\bar{X} := \frac{1}{n} (X_1 + \dots + X_n).$$

$$\Rightarrow \Pr(|\bar{X} - \mu| \geq \varepsilon) \leq \frac{\sigma^2}{n \cdot \varepsilon^2}$$
WHP: empirical averages  $\approx$  expectation value

PF:  $E\bar{X} = \mu$  &  $\text{Var}(\bar{X}) = \frac{1}{n^2} \text{Var}(X_1 + \dots + X_n) = \frac{\sigma^2}{n}$ .  $\leadsto$  Chebyshev. □

One last reminder:

$f$  is called convex if  $p \cdot f(x) + (1-p)f(x') \leq f(px + (1-p)x')$   $\forall p \in [0,1]$

  $\exp, x^2, \dots$  Sufficient for convex:  $f'' \geq 0$   
  $\log, \sqrt{x}, \dots$  Concave:  $f'' \leq 0$

Strongly convex/concave if " $=$ " only holds for  $p=0$  or  $p=1$

Jensen's inequality: If  $f$  convex:  $E[f(X)] \geq f(E[X])$

\* If strongly convex/concave: " $=$ " iff  $X$  is constant.

From next week on we will study:

Entropy of  $X \sim P$ :

$$H(X) = H(P) = E\left[\log \frac{1}{P(X)}\right] = \sum_x P(x) \log \frac{1}{P(x)} = -\sum_x P(x) \log P(x)$$

↑  
ALWAYS BASE 2  
0 · log 0 = 0

\*  $0 \leq H(X) \leq \log(\# \text{outcomes})$

$\begin{matrix} \uparrow & \\ = \text{if } X \text{ constant} & \end{matrix}$      $\begin{matrix} \uparrow & \\ = \text{if } X \text{ uniformly random} & \end{matrix}$

$$E\left[\log \frac{1}{P(X)}\right] \leq \log E\left[\frac{1}{P(X)}\right] \text{ by Jensen}$$

\*  $X \sim \text{Bernoulli}(p)$ : binary entropy function

$$H(X) = -p \cdot \log(p) - (1-p) \log(1-p)$$

